

A Matrix Model for Random Nilpotent Groups

Kelly Delp¹, Tullia Dymarz² and Anshel Schaffer-Cohen³

¹Cornell University ²University of Wisconsin - Madison ³Universidad Mayor de San Andrés

Correspondence to be sent to: dymarz@math.wisc.edu

We study random torsion-free nilpotent groups generated by a pair of random words of length ℓ in the standard generating set of $U_n(\mathbb{Z})$. Specifically, we give asymptotic results about the step properties of the group when the lengths of the generating words are functions of n . We show that the threshold function for asymptotic abelianness is $\ell = c\sqrt{n}$, for which the probability approaches e^{-2c^2} , and also that the threshold function for having full-step, the same step as $U_n(\mathbb{Z})$, is between cn^2 and cn^3 .

1 Introduction

The goal of this paper is to study random finitely-generated torsion-free nilpotent groups (also known as T -groups [2]). Recall that a nilpotent group N is one for which the lower central series eventually terminates:

$$N = N_0 \geq N_1 \geq \dots \geq N_r = \{0\}$$

where $N_i = [N, N_{i-1}]$ is the i th commutator subgroup (i.e. the subgroups generated by commutators of elements in N and N_{i-1}). If r is the first index with $N_r = \{0\}$ then we say that N is nilpotent of *step* r . For more background on nilpotent groups see [6].

Our motivation for studying random nilpotent groups comes from Gromov's study of finitely generated random groups via random presentations (see [7] for a detailed introduction). Roughly speaking Gromov considers groups G_ℓ given by a presentation $G_\ell = \langle S \mid R_\ell \rangle$, where the generating set S is fixed and finite, and the relator set R_ℓ contains a subset of all possible relators of length at most ℓ . A random group is said to have a property P if the probability that G_ℓ has P goes to one as ℓ goes to infinity. Generally the size of R_ℓ depends on ℓ and a chosen density constant $d \in [0, 1]$ where R_ℓ at density d contains on order of the d th power of possible relations of size less than ℓ . Changing d changes the properties of the random group. A fundamental result of Gromov's shows that when the density is greater than $1/2$ the resulting random group is trivial, and when the density is less than $1/2$ then the random group is a so-called *hyperbolic* group. Unfortunately, nilpotent groups are not hyperbolic so this model is unsatisfactory for studying random nilpotent groups. For a recent generalization of Gromov's idea to quotients of free nilpotent groups see [1].

The model we study is motivated by a well-known theorem [4] which states that any finitely-generated, torsion-free nilpotent group appears as a subgroup of $U_n(\mathbb{Z})$, the group of $n \times n$ upper-triangular matrices with ones on the diagonal and entries in \mathbb{Z} .

Let $E_{i,j}$ be the elementary matrix that differs from the identity matrix I_n by containing a one at position (i, j) and set $A_i = E_{i,i+1}$. Then the set $S = \{A_1^{\pm 1}, \dots, A_{n-1}^{\pm 1}\}$ of *superdiagonal elementary matrices* is the *standard generating set* for $U_n(\mathbb{Z})$. Our random subgroups will be generated by taking two simple random walks of length ℓ on the Cayley graph of $U_n(\mathbb{Z})$ given by the generating set S . These two random walks define two words, V, W that generate a subgroup

$$G_{\ell,n} := \langle V, W \rangle \leq U_n(\mathbb{Z}).$$

We are interested in the asymptotic properties of $G_{\ell,n}$ as $\ell \rightarrow \infty$. For example, when n is fixed one can show that the probability that $G_{\ell,n}$ is abelian goes to zero as $\ell \rightarrow \infty$. If ℓ is a function of n , then the asymptotic abelianness depends on the rate of growth.

Before giving the precise statement of our results, we recall the Landau notation that we use to describe the growth rate of ℓ :

- If $f(n) \in O(g(n))$ then there exist numbers c and N , so that $n > N$ implies $f(n) < cg(n)$.
- If $f(n) \in o(g(n))$ then for all $c > 0$, there exists an N , so that $n > N$ implies $f(n) < cg(n)$.
- If $f(n) \in \omega(g(n))$ then for all $c > 0$, there exists an N , so that $n > N$ implies $f(n) > cg(n)$.
- We write $f(n) \sim g(n)$ if $\lim_{n \rightarrow \infty} f(n)/g(n) = 1$.

Let P be a property of a group. For a particular length function $\ell(n)$, we say $G_{\ell,n}$ is *asymptotically almost surely* (a.a.s) P if the probability that $G_{\ell,n}$ has P approaches 1 as n approaches infinity. In Section 4 we prove the following theorem:

Theorem 1.1. Let $G_{\ell,n}$ be a subgroup of $U_n(\mathbb{Z})$ generated by two random walks of length ℓ in the standard generating set S and suppose ℓ is a function of n .

1. If $\ell \in o(\sqrt{n})$ then asymptotically almost surely $G_{\ell,n}$ is abelian.
2. If $\ell = c\sqrt{n}$ then the probability that $G_{\ell,n}$ is abelian approaches e^{-2c^2} as $n \rightarrow \infty$.
3. If $\ell \in \omega(\sqrt{n})$, then asymptotically almost surely $G_{\ell,n}$ is not abelian.

□

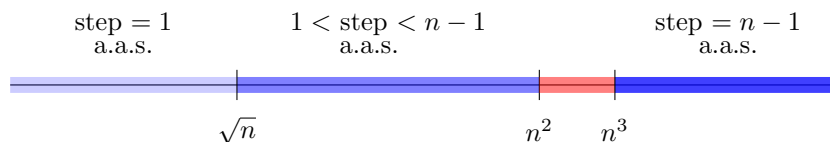
Another property we focus on in this paper is the step of $G_{\ell,n}$. Note that $U_n(\mathbb{Z})$ is a step $n - 1$ nilpotent group. We say that $G_{\ell,n}$ has *full step* if it is also of step $n - 1$. We show that the threshold function for being full step lies between n^2 and n^3 .

Theorem 1.2. Let $G_{\ell,n}$ be a subgroup of $U_n(\mathbb{Z})$ generated by two random walks of length ℓ in the standard generating set S and suppose ℓ is a function of n .

1. If $\ell \in o(n^2)$ then asymptotically almost surely $G_{\ell,n}$ does not have full step.
2. If $\ell \in \omega(n^3)$ then asymptotically almost surely $G_{\ell,n}$ has full step.

□

Theorem 1.2 is proven in Section 5. These theorems are summarized by the following diagram.



1.1 Outline

As random walks, V, W are given by $V = V_1 V_2 \cdots V_\ell$ and $W = W_1 W_2 \cdots W_\ell$ where $V_i, W_i \in S$. To prove Theorem 1.1, we define a sufficient condition for commuting, called *supercommuting*.

Definition 1.3. Let $V = V_1 V_2 \cdots V_\ell$ and $W = W_1 W_2 \cdots W_\ell$ where V_i and W_i are elements in the U_n generating set S . The words V and W **supercommute** if every V_i commutes with every W_j . □

We show that when $\ell \in o(n)$, supercommuting and commuting are asymptotically equivalent, and that the threshold for supercommuting is at $\ell = c\sqrt{n}$. For Theorem 1.2 most of the results are a matter of analyzing the entries on the superdiagonals of our generators V and W . The $(i, i + 1)$ superdiagonal entry of V , which we denote by v_i , is the sum over the number of $A_i^{\pm 1}$ that occur in the walk, where A_i contributes $+1$, and its inverse -1 . Therefore the vector of superdiagonal entries is the endpoint of a random walk in \mathbb{Z}^{n-1} ; while these are well studied objects, most of the study has been on walks in a fixed dimension n . In our case, both the dimension n , and the length of the walk are going to ∞ . We gather these results in Section 3.

2 Preliminaries

Many of the results in this paper depend on the superdiagonal entries $v_{i,i+1}$ and $w_{i,i+1}$ of V and W . For this reason we adopt the shorthand $z_i := z_{i,i+1}$ for any matrix Z .

The following proposition gives a necessary condition for commuting in U_n .

Lemma 2.1. Let $W = [w_{i,j}]$ and $V = [v_{i,j}]$ be matrices in U_n . If W and V commute then $w_i v_{i+1} = w_{i+1} v_i$ for all $1 \leq i \leq n - 2$. □

Proof. This is a straightforward computation. The first superdiagonal of $C = VWV^{-1}W^{-1}$ vanishes and the second superdiagonal entries are given by $c_{i,i+2} = w_{i+1}v_i - w_i v_{i+1}$. ■

Corollary 2.2. The elementary superdiagonal matrices $A_i^{\pm 1}$, $A_j^{\pm 1}$ commute if and only if $|i - j| \neq 1$. □

Next we study the k th commutator subgroup of $G_{\ell,n} = \langle V, W \rangle$. Note that in a nilpotent group the k th commutator subgroup is generated by all m -fold commutators for $m \geq k$ of the form

$$[B_1[B_2 \cdots [B_m, B_{m+1}]]]$$

where the B_i are chosen from a fixed generating set (see for example Lemma 1.7 in [4]). Therefore to test that $G_{\ell,n}$ is k -step nilpotent we only need to check that $[B_1[B_2 \cdots [B_k, B_{k+1}]]] = I$ when $B_i \in \{V, W\}$.

In Lemma 2.1 we noted that taking a commutator resulted in a matrix with zeros along the first superdiagonal. In the next lemma we show that taking a k^{th} commutator results in zeros on the first k superdiagonals. We also give a recursive formula for the entries on the $(k + 1)^{\text{st}}$ superdiagonal using iterated two dimensional determinants.

Lemma 2.3. Let $C^k = [c_{i,j}^k]$ be a k -fold commutator of two matrices V, W ; then $c_{i,j}^k = 0$ when $i < j \leq i + k$ and

$$c_{i,k+i+1}^k = \det \begin{bmatrix} z_{i,i+1} & c_{i,k+i}^{k-1} \\ z_{k+i,k+i+1} & c_{i+1,k+i+1}^{k-1} \end{bmatrix} \quad (1)$$

where $Z = [z_{i,j}]$ and either $Z = V$ or $Z = W$. □

Proof. We prove this result by induction, where the base case is given in the proof of Lemma 2.1. Assume C^{k-1} is given, and for convenience let $K = C^{k-1}$. Since the first $k - 1$ superdiagonals of K contain all zeros, computing $C^k = ZKZ^{-1}K^{-1}$ yields zeros on the first k superdiagonals, and on the $(i, i + k + 1)$ -diagonal we have

$$z_{i,i+1}c_{i+1,k+i+1}^{k-1} - z_{k+i,k+i+1}c_{i,k+i}^{k-1}.$$

To help see this, note that when the first nonzero superdiagonals of Z, C^{k-1}, C^k are overlaid the resulting matrix is the following.

$$\begin{pmatrix} \ddots & \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ & 1 & z_{i,i+1} & \cdots & c_{i,k+i}^{k-1} & c_{i,i+k+1}^k & \cdots & \cdots \\ & & 1 & \cdots & \vdots & c_{i+1,k+i+1}^{k-1} & \cdots & \cdots \\ & & & \ddots & \vdots & \vdots & \cdots & \cdots \\ & & & & 1 & z_{k+i,k+i+1} & \cdots & \cdots \\ & & & & & 1 & \cdots & \cdots \\ & & & & & & \ddots & \ddots \end{pmatrix}$$

Our first application of Lemma 2.3 is the following lemma, which shows that $G_{\ell,n}$ cannot be full step if V, W have a matching pair of zeros on their superdiagonals.

Lemma 2.4. Given $G_{\ell,n} = \langle V, W \rangle$, if there is some $1 \leq d \leq n - 1$ such that $v_d = w_d = 0$ then the step of $G_{\ell,n}$ is bounded by $\max\{d - 1, n - 1 - d\}$. □

Proof. Recall that $v_d = v_{d,d+1}$ and similarly $w_d = w_{d,d+1}$. By Lemma 2.3 we have that for $C^1 = [V, W]$

$$c_{d-1,d+1}^1 = \det \begin{bmatrix} v_{d-1,d} & w_{d-1,d} \\ v_{d,d+1} & w_{d,d+1} \end{bmatrix} = 0$$

since the bottom row of this two by two matrix has both entries to zero. Similarly

$$c_{d,d+2}^1 = \det \begin{bmatrix} v_{d,d+1} & w_{d,d+1} \\ v_{d+1,d+2} & w_{d+1,d+2} \end{bmatrix} = 0$$

since the top row of the two by two matrix has entries both equal to zero. Inductively, by Equation 1, we have that $c_{d-k,d+1}^k = c_{d,d+k+1}^k = 0$. This is because either the top or bottom row of the matrix in Equation 1 will have both entries equal to zero. Alternatively, if both $c_{i,k+i}^{k-1}$ and $c_{i+1,k+i+1}^{k-1}$ are zero then $c_{i,k+i+1}^k = 0$ since then the righthand column of the matrix in Equation 1 will have both entries zero. In particular, $c_{i,j}^k$ is zero if $k > \max\{d - 1, n - 1 - d\}$. ■

Lemma 2.3 also leads us to define a modified determinant product which gives us a method to calculate the entries of the first nonzero superdiagonal of a iterated commutator product of upper triangular matrices given their first superdiagonal entries.

Definition 2.5. Let $\vec{a} = (a_1, \dots, a_s)$ and $\vec{b} = (b_1, \dots, b_m)$ be vectors with $s \geq m$ and set $s - m = p$; then $[\vec{a} \ \vec{b}]$ is the $m - 1$ dimensional vector given by

$$[\vec{a} \ \vec{b}] := \left(\left| \begin{array}{cc} a_1 & b_1 \\ a_{p+2} & b_2 \end{array} \right|, \left| \begin{array}{cc} a_2 & b_2 \\ a_{p+3} & b_3 \end{array} \right|, \dots, \left| \begin{array}{cc} a_{m-1} & b_{m-1} \\ a_{p+m} & b_m \end{array} \right| \right).$$

□

Lemma 2.6. Let \vec{b}_i be the vector containing the $n - 1$ main superdiagonal entries of an $n \times n$ unipotent matrix B_i labeled from top left to bottom right. Then the $(k + 1)^{\text{st}}$ superdiagonal entries of the k -fold commutator $[B_1[B_2 \cdots [B_k, B_{k+1}]]]$ are given by the $(n - k)$ dimensional vector

$$[\vec{b}_1[\vec{b}_2 \cdots [\vec{b}_k \ \vec{b}_{k+1}] \cdots]].$$

□

This lemma can be proved by direct computation or by inspecting the proof of Lemma 2.3. To illustrate this result, consider the following examples, the second of which will be used in Section 5.

Example 2.7. We consider the commutator $[D, [C, [A, B]]]$ where the superdiagonal entries of A are given by (a_1, \dots, a_{n-1}) and similarly for B, C, D . The first three superdiagonals are all zero while the fourth superdiagonal has entries given by

$$\left(\begin{array}{c} d_1 \\ \\ \\ d_4 \end{array} \left| \begin{array}{c} c_1 \\ \\ c_3 \\ \\ c_2 \\ \\ c_4 \end{array} \left| \begin{array}{cc} a_1 & b_1 \\ a_2 & b_2 \\ a_3 & b_3 \\ a_2 & b_2 \\ a_3 & b_3 \\ a_3 & b_3 \\ a_4 & b_4 \end{array} \right| \right| \dots \left| \begin{array}{c} d_{n-4} \\ \\ \\ d_{n-1} \end{array} \left| \begin{array}{c} c_{n-4} \\ \\ c_{n-2} \\ \\ c_{n-3} \\ \\ c_{n-1} \end{array} \left| \begin{array}{cc} a_{n-4} & b_{n-4} \\ a_{n-3} & b_{n-3} \\ a_{n-3} & b_{n-3} \\ a_{n-2} & b_{n-2} \\ a_{n-3} & b_{n-3} \\ a_{n-2} & b_{n-2} \\ a_{n-2} & b_{n-2} \\ a_{n-1} & b_{n-1} \end{array} \right| \right| \right).$$

□

Example 2.8. Consider the commutator

$$\underbrace{[W, [W, \dots [W, V]]]}_{n-2}$$

where V, W are $n \times n$ upper triangular matrices with main superdiagonals given by the vectors (v_1, \dots, v_{n-1}) and (w_1, \dots, w_{n-1}) respectively. Using the iterated determinant formula we see that the first nonzero superdiagonal has only one entry and is given by

$$K_1 v_1 w_2 w_3 \cdots w_{n-1} + K_2 w_1 v_2 w_3 \cdots w_{n-1} + \cdots + K_{n-1} w_1 \cdots w_{n-2} v_{n-1}$$

where each $K_i = \binom{n-1}{i}$ with alternating signs. □

3 Distribution of the Superdiagonal Entries

In this section we examine the probability of finding zeroes on the superdiagonals of V and W when $\ell \in \omega(n)$. In order to emphasize the dependence on ℓ we write V^ℓ, W^ℓ instead of V, W and v_k^ℓ, w_k^ℓ instead of v_k, w_k for the superdiagonal entries. If we fix n and k we can model v_k^ℓ as the endpoint of a lazy random walk in \mathbb{Z} :

$$v_k^\ell = \sum_{j=1}^{\ell} x_j$$

where $x_j = \pm 1$ with probability $1/2n$ each and $x_j = 0$ with probability $(n - 1)/n$. Likewise for any two $k_1 \neq k_2$ we have an induced lazy random walk on \mathbb{Z}^2 :

$$\begin{pmatrix} v_{k_1}^\ell \\ v_{k_2}^\ell \end{pmatrix} = \sum_{j=1}^{\ell} \begin{pmatrix} x_j \\ y_j \end{pmatrix}$$

where $(x_j, y_j) = (\pm 1, 0)$ or $(0, \pm 1)$ with probability $1/2n$ each, and $(x_j, y_j) = (0, 0)$ with probability $(n - 2)/n$.

Our goal is to estimate $P(v_k^\ell = 0)$ and $P(v_{k_1}^\ell = v_{k_2}^\ell = 0)$. The proofs of the following lemmas follow the standard proofs of the local central limit theorem for lazy random walks on \mathbb{Z}^d where special attention is paid to the dependence of the estimates on n . (See for example Section 2.3 in [5]). We reproduce them here because we were not able to find this exact formulation in the literature. Morally we rewrite everything in terms of $\lambda = \ell/n$ and provide error estimates. We can do this as long as $\lambda \rightarrow \infty$ —that is, when $\ell \in \omega(n)$. To make the results in this section more applicable later on, we define a constant $K = 1/\sqrt{2\pi}$.

Lemma 3.1. Suppose $\ell \in \omega(n)$. Then for a fixed $1 \leq k \leq n$ we have

$$P(v_k^\ell = 0) \sim K \sqrt{\frac{n}{\ell}}$$

□

Proof. We begin by noting that the characteristic function of x_j is given by

$$\phi(t) = E(e^{itx_j}) = 1 - \frac{1}{n} + \frac{1}{2n}(e^{it} + e^{-it}) = 1 - \frac{1}{n}(1 - \cos t)$$

and the characteristic function of v_k^ℓ which is

$$\phi(t)^\ell = \left(1 - \frac{1}{n}(1 - \cos t)\right)^\ell.$$

Therefore

$$P(v_k^\ell = 0) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \left(1 - \frac{1}{n}(1 - \cos t)\right)^\ell dt.$$

The methods used to estimate this integral are identical to the ones used in the more general proof of Lemma 3.3 below so we do not produce them here. The above integral is transformed to

$$P(v_k^\ell = 0) = \frac{\sqrt{\frac{n}{\ell}}}{\sqrt{2\pi}} \left(\int_{\mathbb{R}} e^{-s^2/2} ds + o(1) \right)$$

■

Since v_k^ℓ and w_k^ℓ are independent we have the following corollary:

Corollary 3.2. Suppose $\ell \in \omega(n)$. For fixed k , $P(v_k = w_k = 0) \sim K^2 n/\ell$.

□

Next we prove an estimate on the probability of having a pair of zeros in fixed coordinates $k_1 \neq k_2$.

Lemma 3.3. Suppose $\ell \in \omega(n)$. Then for fixed $k_1 \neq k_2$,

$$P(v_{k_1}^\ell = v_{k_2}^\ell = 0) \sim K^2 \frac{n}{\ell}.$$

□

Proof. We begin by computing the characteristic function of (x_j, y_j) which is given by

$$\phi(t_1, t_2) = E(e^{i(t_1 x_j + t_2 y_j)}) = 1 - \frac{1}{n}(1 - \cos t_1) - \frac{1}{n}(1 - \cos t_2)$$

and the characteristic function of $(v_{k_1}^\ell, v_{k_2}^\ell) = (\sum_{j=1}^\ell x_j, \sum_{j=1}^\ell y_j)$ which is

$$\phi(t)^\ell = \left(1 - \frac{1}{n}(1 - \cos t_1) - \frac{1}{n}(1 - \cos t_2)\right)^\ell.$$

Therefore

$$\begin{aligned} P(v_{k_1}^\ell = v_{k_2}^\ell = 0) &= \frac{1}{(2\pi)^2} \iint_{[-\pi, \pi]^2} \left(1 - \frac{1}{n}(1 - \cos t_1) - \frac{1}{n}(1 - \cos t_2)\right)^\ell dt_1 dt_2. \\ &= \frac{1}{(2\pi)^2} \iint_{[-\pi, \pi]^2} \left(1 - \frac{|\theta|^2}{2n} + \frac{1}{n}h(\theta)\right)^\ell d\theta \end{aligned}$$

where $\theta = (t_1, t_2)$ and $h(\theta) = \sum_{i=2}^\infty (-1)^i \frac{t_1^{2i} + t_2^{2i}}{(2i)!} \in O(|\theta|^4)$.

We use the Taylor expansion $\log(1+x) = \sum_{i=1}^\infty \frac{x^i}{i}$ that is valid for $|x| \leq 1$ to write

$$\log(\phi(\theta)) = \log \left(\underbrace{1 - \frac{|\theta|^2}{2n} + \frac{1}{n}h(\theta)}_x \right) = -\frac{|\theta|^2}{2n} + \frac{1}{n}h(\theta) + f(\theta, 1/n) \quad (2)$$

where

$$f(\theta, 1/n) = \sum_{j=2}^\infty \frac{1}{j} \left(\frac{1}{n} \sum_{i=1}^\infty (-1)^i \frac{t_1^{2i} + t_2^{2i}}{(2i)!} \right)^j = O(|\theta|^4).$$

This expansion is valid for

$$\left| -\frac{|\theta|^2}{2n} + \frac{1}{n}h(\theta) \right| = \frac{1}{n} \left| -\frac{|\theta|^2}{2} + h(\theta) \right| \leq 1$$

which holds as long as $|\theta| < \delta$ where δ does not depend on n . (It holds for $n = 1$ and so it holds for all n). Let $\lambda = \ell/n$. Now use a change of variable $\theta = s/\sqrt{\lambda} = s\sqrt{\frac{n}{\ell}}$ in Equation 2 and multiply both sides by ℓ to get

$$\ell \log \left(\phi \left(s/\sqrt{\lambda} \right) \right) = -\frac{|s|^2}{2} + \underbrace{\frac{\ell}{n} h(s/\sqrt{\lambda}) + \bar{f}(s, 1/\ell, \lambda)}_{g_n(\ell, s)}$$

where $\bar{f} = \ell f$ is given by

$$\bar{f}(s, 1/\ell, \lambda) = \sum_{j=2}^\infty \frac{1}{j \ell^{j-1}} \left(\sum_{i=1}^\infty (-1)^i \frac{1}{\lambda^{i-1}} \frac{s_1^{2i} + s_2^{2i}}{(2i)!} \right)^j.$$

This expansion is valid as long as $|s| \leq \delta\sqrt{\lambda}$. Note that when $n = 1$ we have

$$\bar{f}(s, 1/\ell, \ell) = \sum_{j=2}^\infty \frac{1}{j \ell^{j-1}} \left(\sum_{i=1}^\infty (-1)^i \frac{1}{\ell^{i-1}} \frac{s_1^{2i} + s_2^{2i}}{(2i)!} \right)^j$$

and since $\ell > \lambda$ we have that

$$\sum_{j=2}^\infty \frac{1}{j \ell^{j-1}} \left(\sum_{i=1}^\infty \frac{1}{\lambda^{i-1}} \frac{s_1^{2i} + s_2^{2i}}{(2i)!} \right)^j \leq \sum_{j=2}^\infty \frac{1}{j \lambda^{j-1}} \left(\sum_{i=1}^\infty \frac{1}{\lambda^{i-1}} \frac{s_1^{2i} + s_2^{2i}}{(2i)!} \right)^j.$$

Then

$$|g_n(\ell, s)| \leq \lambda |h(s/\sqrt{\lambda})| + |\bar{f}(s, 1/\ell, \lambda)| \leq \lambda |h(s/\sqrt{\lambda})| + \frac{c|s|^4}{\lambda}$$

where c can be chosen independent of n . Note that

$$\lambda h(s/\sqrt{\lambda}) = \lambda \sum_{i=2}^\infty (-1)^i \frac{1}{\lambda^i} \frac{s_1^{2i} + s_2^{2i}}{(2i)!} = \sum_{i=2}^\infty (-1)^i \frac{1}{\lambda^{i-1}} \frac{s_1^{2i} + s_2^{2i}}{(2i)!}$$

so $\lambda |h(s/\sqrt{\lambda})| = o(|s|^2)$ and so we can find $0 < \epsilon \leq \delta$ such that for $|s| \leq \epsilon\sqrt{\lambda}$

$$|g_n(s, \ell)| \leq \frac{|s|^2}{4}.$$

Let $F_{\ell, n}(s) = e^{g_n(\ell, s)} - 1$ and let

$$\bar{p}_\ell(0) = \frac{1}{(2\pi)^2 \lambda} \int_{\mathbb{R}^2} e^{-\frac{|s|^2}{2}} ds = \frac{1}{2\pi\lambda}$$

be the integral of a two-variable standard normal distribution (see for example Equation 2.2 in [5]). Then

$$\begin{aligned} P(v_{k_1}^\ell = v_{k_2}^\ell = 0) &= \frac{1}{(2\pi)^2} \iint_{[-\pi, \pi]^2} \phi(\theta)^\ell d\theta \\ &= \frac{1}{(2\pi)^2} \iint_{[-\pi, \pi]^2} \left(1 - \frac{|\theta|^2}{2n} + \frac{1}{n} h(\theta)\right)^\ell d\theta \\ &= \frac{1}{(2\pi)^2 \lambda} \iint_{[-\pi\sqrt{\lambda}, \pi\sqrt{\lambda}]^2} e^{-|s|^2/2} (F_{\ell, n}(s) + 1) ds \\ &= \frac{1}{(2\pi)^2 \lambda} \left(A_n(\epsilon, \ell) + \iint_{|s| \leq \epsilon\sqrt{\lambda}} e^{-|s|^2/2} (F_{\ell, n}(s) + 1) ds \right) \\ &= \bar{p}_\ell(0) + B_n(\epsilon, \ell) + \frac{1}{(2\pi)^2 \lambda} \left(A_n(\epsilon, \ell) + \iint_{|s| \leq \epsilon\sqrt{\lambda}} e^{-\frac{|s|^2}{2}} F_{\ell, n}(s) ds \right) \end{aligned}$$

where

$$|A_n(\epsilon, \ell)| = \left| \iint_{[-\pi\sqrt{\lambda}, \pi\sqrt{\lambda}]^2 \setminus \{|s| \leq \epsilon\sqrt{\lambda}\}} \phi(s/\sqrt{\lambda})^\ell ds \right| \leq C\lambda e^{-\beta\lambda}$$

where C and β do not depend on n since $|\phi(\theta)| \leq 1 - \frac{b}{n}|\theta|^2 \leq e^{-\frac{b}{n}|\theta|^2}$ (where b does not depend on n) for all $\theta \in [-\pi, \pi]^2$ and so for $|s| \geq \epsilon\sqrt{\lambda}$ we have $\phi(s/\sqrt{\lambda}) \leq e^{-\beta/n}$. Likewise

$$|B_n(\epsilon, \ell)| = \left| \frac{1}{(2\pi)^2 \lambda} \iint_{|s| > \epsilon\sqrt{\lambda}} e^{-|s|^2/2} ds \right| \leq C' e^{-\beta'\lambda}$$

where β' and C' do not depend on n . Finally as long as $|s| \leq \lambda^{\frac{1}{8}}$ we have

$$|F_{\ell, n}(s)| \leq |e^{g_n(\ell, s)} - 1| \leq C'' g_n(\ell, s) \leq \frac{C'' |s|^4}{\lambda}$$

where C'' does not depend on n . Therefore we have

$$\left| \iint_{|s| \leq \lambda^{1/8}} e^{-\frac{|s|^2}{2}} F_{\ell, n}(s) ds \right| \leq \frac{C''}{\lambda} \int_{\mathbb{R}^2} |s|^4 e^{-\frac{|s|^2}{2}} ds \leq \frac{C'''}{\lambda}.$$

This leaves us only to estimate the integral for $\lambda^{1/8} \leq |s| \leq \epsilon\sqrt{\lambda}$ where we have the bound $|F_{\ell, n}(s)| \leq e^{-\frac{|s|^2}{4}} + 1$. The integral then can be estimated as follows

$$\left| \iint_{\lambda^{1/8} \leq |s| \leq \epsilon\sqrt{\lambda}} e^{-\frac{|s|^2}{2}} F_{\ell, n}(s) ds \right| \leq 2 \iint_{|s| \geq \lambda^{1/8}} e^{-\frac{|s|^2}{4}} ds \leq \bar{C} e^{-\zeta\lambda^{1/4}}.$$

This gives the desired result. ■

Corollary 3.4. Suppose $\ell \in \omega(n)$. For fixed $k_1 \neq k_2$,

$$P(v_{k_1}^\ell = v_{k_2}^\ell = w_{k_1}^\ell = w_{k_2}^\ell = 0) \sim K^4 \left(\frac{n}{\ell}\right)^2.$$

□

Proof. This follows from Lemma 3.3 and the fact that

$$P(v_{k_1}^\ell = v_{k_2}^\ell = w_{k_1}^\ell = w_{k_2}^\ell = 0) = P(v_{k_1}^\ell = v_{k_2}^\ell = 0) P(w_{k_1}^\ell = w_{k_2}^\ell = 0).$$

■

Lemma 3.5. Suppose $\ell \in \omega(n)$ and suppose $a_i = a_i(\ell)$ for $1 \leq i \leq n-1$, with $P(a_1 \neq 0) \rightarrow 1$ as $\ell \rightarrow \infty$. Then $P(a_1 v_1 + a_2 v_2 + \cdots + a_{n-1} v_{n-1} = 0) \rightarrow 0$ as $\ell \rightarrow \infty$. \square

Proof.

$$\begin{aligned} P\left(\sum_{i=1}^n a_i v_i = 0\right) &= P\left(v_1 = -\sum_{i=2}^n \frac{a_i}{a_1} v_i = 0 \mid a_1 \neq 0\right) P(a_1 \neq 0) \\ &\quad + P\left(\sum_{i=1}^n a_i v_i = 0 \mid a_1 = 0\right) P(a_1 = 0) \\ &\leq P\left(v_1 = -\sum_{i=2}^n \frac{a_i}{a_1} v_i = 0 \mid a_1 \neq 0\right) + P(a_1 = 0) \\ &\leq P(v_1 = 0) + P(a_1 = 0) \end{aligned}$$

since the most likely value for v_1 is 0 and therefore by Lemma 3.1 this limit goes to zero. \blacksquare

4 Asymptotic Abelianess

In this section we prove Theorem 1.1. To check that $G_{\ell,n}$ is abelian we only need to check that V, W commute. Most of our analysis involves the notion of *supercommuting* that we defined in the introduction. Recall that for two words $V = V_1 V_2 \cdots V_\ell$ and $W = W_1 W_2 \cdots W_\ell$ with $V_i, W_i \in S$ to supercommute, every V_i must commute with every W_j .

Clearly supercommuting is a sufficient (but not necessary) condition for commuting. However, when $\ell \in o(n)$, the probability of V and W commuting but not supercommuting goes to zero as $n \rightarrow \infty$. Therefore, when ℓ is in this class, these two notions of commuting are asymptotically equivalent. To prove this fact, we begin by defining the function

$$\sigma_i(Z) := \begin{cases} 1 & \text{if } Z = A_i \\ -1 & \text{if } Z = A_i^{-1} \\ 0 & \text{otherwise.} \end{cases}$$

Since multiplication in U_n is additive on the superdiagonal elements,

$$v_i = \sum_{j=1}^{\ell} \sigma_i(V_j) \quad w_i = \sum_{j=1}^{\ell} \sigma_i(W_j).$$

In other words, the i^{th} superdiagonal entry of V is a count of the number of times one of $A_i^{\pm 1}$ appears in the word $V = V_1 \dots V_n$, where A_i contributes $+1$, and its inverse -1 . Since ℓ is growing more slowly than the size of our matrix (and hence more slowly than the size of our generating set S), the probability of seeing a particular A_i in an ℓ -step walk approaches zero. We make this precise in the following lemma.

Lemma 4.1. Suppose $\ell \in o(n)$. For fixed $1 \leq i \leq n-1$ and $Z = Z_1 Z_2 \cdots Z_\ell$, where $Z_i \in S = \{A_1^{\pm 1}, \dots, A_{n-1}^{\pm 1}\}$,

$$P(\sigma_i(Z_j) \neq 0 \text{ for some } 1 \leq j \leq \ell) \rightarrow 0$$

as $n \rightarrow \infty$. \square

Proof. For fixed j ,

$$P(\sigma_i(Z_j) = 0) = \left(1 - \frac{2}{2(n-1)}\right) = \left(1 - \frac{1}{n-1}\right).$$

Since the Z_j 's are independent,

$$P(\sigma_i(Z_j) = 0 \text{ for all } j) = \left(1 - \frac{1}{n-1}\right)^\ell.$$

Since $\ell \in o(n)$, the limit of this probability is 1, and so its negation—the probability that $\sigma_i(Z_j) \neq 0$ for some j —goes to 0. \blacksquare

Now suppose that A_i appears at least once in our word $Z_1 Z_2 \cdots Z_\ell$. Lemma 4.1 implies that it, or its inverse, almost surely does not appear again.

Corollary 4.2. Suppose $\ell \in o(n)$ and $Z = Z_1 Z_2 \cdots Z_\ell$. For a fixed $1 \leq i \leq n-1$, the i^{th} superdiagonal entry z_i of Z satisfies

$$P(z_i = \pm 1 \mid \sigma_i(Z_j) \neq 0 \text{ for some } j) \rightarrow 1$$

as $n \rightarrow \infty$. □

Proof. This follows from the fact that $P(z_i = \pm 1 \mid \sigma_i(Z_j) \neq 0 \text{ for some } j, \text{ and } \sigma_i(Z_k) = 0 \text{ for all } k \neq j) = 1$ and

$$P(\sigma_i(Z_k) = 0 \text{ for all } k \neq j) = \left(1 - \frac{1}{n-1}\right)^{\ell-1} \rightarrow 1$$

as $\ell \rightarrow \infty$. ■

Lemma 4.3. When $\ell \in o(n)$,

$$P(V \text{ and } W \text{ commute but do not supercommute}) \rightarrow 0.$$

□

Proof. Note that

$$\begin{aligned} & P(V \text{ and } W \text{ commute but do not supercommute}) \\ & \leq P(V \text{ and } W \text{ commute} \mid V \text{ and } W \text{ do not supercommute}). \end{aligned}$$

We will call this latter (conditional) event \mathcal{C} and show that $P(\mathcal{C}) \rightarrow 0$.

Let A_i and A_{i+1} be called *neighboring* elementary matrices. If V and W do not supercommute, then Corollary 2.2 implies the words V and W must contain neighboring matrices. Without loss of generality, this implies there must be some $1 < k \leq n-1$ and some $1 \leq i, j \leq \ell$ such that $\sigma_{k-1}(W_i) \neq 0$ and $\sigma_k(V_j) \neq 0$. We bound $P(\mathcal{C})$ above by considering the events $w_{k-1} \neq \pm 1$, $v_k \neq \pm 1$ and the joint event $w_{k-1} = \pm 1, v_k = \pm 1$. While these three events are not mutually exclusive, they do cover all possibilities.

$$\begin{aligned} P(\mathcal{C}) & \leq P(\mathcal{C} \mid w_{k-1} \neq \pm 1) P(w_{k-1} \neq \pm 1) + P(\mathcal{C} \mid v_k \neq \pm 1) P(v_k \neq \pm 1) \\ & \quad + P(\mathcal{C} \mid w_{k-1}, v_k = \pm 1) P(w_{k-1}, v_k = \pm 1). \end{aligned}$$

By Corollary 4.2 the first two terms go to 0 and the last term goes to just $P(\mathcal{C} \mid w_{k-1}, v_k = \pm 1)$. By Lemma 2.1, this is at most

$$\begin{aligned} P(\mathcal{C} \mid w_{k-1}, v_k = \pm 1) & \leq P(w_{k-1} v_k - w_k v_{k-1} = 0 \mid w_{k-1}, v_k = \pm 1) \\ & \leq P(w_k v_{k-1} \neq 0) \\ & \leq P(v_{k-1} \neq 0) \end{aligned}$$

and $P(v_{k-1} \neq 0) \rightarrow 0$ by Lemma 4.1. ■

4.1 Part 1 of Theorem 1.1: when $\ell(n) \in o(\sqrt{n})$.

In this case, we can use a counting argument to show that V and W supercommute.

Lemma 4.4. Assume that $\ell \in o(\sqrt{n})$, $V = V_1 V_2 \cdots V_\ell$, and $W = W_1 W_2 \cdots W_\ell$. Let F be the number of pairs i, j for which V_i and W_j fail to commute. Then the expected value $E(F) \rightarrow 0$ as $n \rightarrow \infty$. □

Proof. Let $\gamma_{i,j}$ be an indicator random variable whose value is 1 precisely when $V_i W_j \neq W_j V_i$. By Corollary 2.2, for each k , there are at most 2 values of i such that V_i does not commute with $A_k^{\pm 1}$. Since $W_j = A_k^{\pm 1}$ for some $1 \leq k \leq n-1$, when $2 \leq k \leq n-2$, the probability that V_i does not commute with W_j is $\frac{4}{2(n-1)} = \frac{2}{n-1}$; when k is equal to 1 or $n-1$, the probability is $\frac{2}{2(n-1)} = \frac{1}{n-1}$. Therefore the probability $P(V_i W_j \neq W_j V_i) \leq \frac{2}{n-1}$ for all i and j . Since F counts the number of non-commuting pairs V_i, W_j , we have

$$F = \sum_{i=1}^{\ell} \sum_{j=1}^{\ell} \gamma_{i,j}.$$

By linearity of expected value,

$$\begin{aligned}
\mathbb{E}(F) &= \sum_{i=1}^{\ell} \sum_{j=1}^{\ell} \mathbb{E}(\gamma_{i,j}). \\
&= \sum_{i=1}^{\ell} \sum_{j=1}^{\ell} \mathbb{P}(V_i W_j \neq W_j V_i) \\
&\leq \sum_{i=1}^{\ell} \sum_{j=1}^{\ell} \frac{2}{n-1} \\
&\leq \ell^2 \left(\frac{2}{n-1} \right).
\end{aligned}$$

Since $\ell \in o(\sqrt{n})$ then $\ell^2 \in o(n)$ and

$$\lim_{n \rightarrow \infty} \mathbb{E}(F) \leq \lim_{n \rightarrow \infty} \frac{2\ell^2}{n-1} = 0. \quad \blacksquare$$

Corollary 4.5. If $\ell \in o(\sqrt{n})$ then V and W supercommute asymptotically almost surely. \square

Proof. The elements V and W supercommute precisely when every V_i commutes with every W_j , that is when $F = 0$. Since F is a nonnegative integer random variable, and $\mathbb{E}(F) \rightarrow 0$ we have that $\mathbb{P}(F = 0) \rightarrow 1$. \blacksquare

4.2 Part 2 of Theorem 1.1: when $\ell = c\sqrt{n}$.

We start with a heuristic argument. For V and W to supercommute, V_i must commute with W_j for all $1 \leq i, j \leq \ell$. The probability that a given V_i and W_j commute is $1 - 2/(n-1)$ for most cases. Since there are ℓ^2 such pairs, the probability that they all commute is

$$\left(1 - \frac{2}{n-1}\right)^{\ell^2} = \left(1 - \frac{2}{n-1}\right)^{c^2 n} \rightarrow \frac{1}{e^{2c^2}}.$$

This argument assumes independence of each V_i, W_j pair commuting, which does not in general hold. However, we are able to show that limiting probability for abelianess is nonetheless $1/e^{2c^2}$, as predicted.

If we fix the V_i 's, there is a specific set of k 's for which $A_k^{\pm 1}$ fails to commute with at least one V_i . Let B be the number of such k 's; then since the W_j 's are chosen independently, the probability that all of them commute with V is given by

$$\mathbb{P}(V \text{ and } W \text{ supercommute}) = \left(1 - \frac{B}{n-1}\right)^{\ell}. \quad (3)$$

Now we have to say something about the distribution of B . Imagine a row of $n-1$ bins. For each $V_i = A_k^{\pm 1}$, we put a ball in bin $k-1$ and a ball in bin $k+1$. Then B is the number of non-empty bins. Since there are 2ℓ balls, two* for each V_i , we have $2 \leq B \leq 2\ell$. Let D be the difference $2\ell - B$. We will show that the expected value of D approaches a constant.

Lemma 4.6. If $\ell = c\sqrt{n}$ then $\mathbb{E}(D) \rightarrow 2c^2$ as $n \rightarrow \infty$. \square

Proof. Let $V = V_1 V_2 \cdots V_{\ell}$. First, we count the number X of ‘‘empty bins’’. We write $X = \sum X_i$, where

$$X_i = \begin{cases} 0 & \text{if } A_{i+1} \text{ or } A_{i-1} \text{ appears in the word } V \\ 1 & \text{otherwise.} \end{cases}$$

Note that the behaviors for the end bins (when $i = 1$ or $i = n-1$) are slightly different than the other bins but asymptotically this difference will not be important. Since each element V_i is chosen independently, we have,

$$\mathbb{E}(X_i) = \mathbb{P}(X_i = 1) = \left(1 - \frac{2}{n-1}\right)^{\ell}.$$

*When $V_i = A_1^{\pm 1}$ or $A_{n-1}^{\pm 1}$ only one ball is added; but this almost never happens as $n \rightarrow \infty$.

Therefore, $E(X) = (n-1) \left(1 - \frac{2}{n-1}\right)^\ell$. Since B is the number of nonempty bins, $B + X = n-1$, and we have,

$$E(B) = (n-1) - (n-1) \left(1 - \frac{2}{n-1}\right)^\ell = (n-1) \left(1 - \left(1 - \frac{2}{n-1}\right)^\ell\right).$$

Finally, since D is the difference $2\ell - B$, the expected value of D is $E(D) = 2\ell - E(B)$. Taking the limit as n goes to infinity gives the result. \blacksquare

In order to evaluate the limit of Equation (3) as $\ell \rightarrow \infty$ we need to control the size of $B = 2\ell - D$. For this we consider two cases: when $D \geq \log \ell$ and when $D \leq \log \ell$.

Lemma 4.7. If $\ell = c\sqrt{n}$ then $P(D \geq \log \ell) \rightarrow 0$ as $n \rightarrow \infty$. \square

Proof. Markov's inequality tells us that $P(D \geq \log \ell) \leq E(D)/\log \ell$. Since $E(D)$ converges to a constant by Lemma 4.6 but $\log \ell$ grows without bound, this probability goes to 0. \blacksquare

Lemma 4.8. If $\ell = c\sqrt{n}$ then $P(G_{\ell,n} \text{ is abelian} \mid D < \log \ell) \rightarrow 1/e^{2c^2}$ as $n \rightarrow \infty$. \square

Proof. Recall (by Lemma 4.3 and Equation 3) that

$$\begin{aligned} \lim_{n \rightarrow \infty} P(G_{\ell,n} \text{ is abelian}) &= \lim_{n \rightarrow \infty} P(W \text{ and } V \text{ supercommute}) \\ &= \lim_{n \rightarrow \infty} \left(1 - \frac{B}{n-1}\right)^\ell \end{aligned}$$

and that by definition of D , $B = 2\ell - D$. Since $0 < D < \log \ell$, we have

$$\left(1 - \frac{2\ell}{n-1}\right)^\ell \leq \left(1 - \frac{2\ell - D}{n-1}\right)^\ell \leq \left(1 - \frac{2\ell - \log \ell}{n-1}\right)^\ell.$$

Using standard techniques (taking the logarithm and using L'Hôpital's rule) one can show that as $n \rightarrow \infty$ both the extreme functions limit to $1/e^{2c^2}$, and the result follows. \blacksquare

Lemma 4.9. If $\ell = c\sqrt{n}$ then $P(G_{\ell,n} \text{ is abelian}) \rightarrow 1/e^{2c^2}$ as $n \rightarrow \infty$. \square

Proof. We have

$$\begin{aligned} \lim_{n \rightarrow \infty} P(G_{\ell,n} \text{ is abelian}) &= \lim_{n \rightarrow \infty} P(G_{\ell,n} \text{ is abelian} \mid D < \log \ell) P(D < \log \ell) \\ &\quad + \lim_{n \rightarrow \infty} P(G_{\ell,n} \text{ is abelian} \mid D \geq \log \ell) P(D \geq \log \ell). \end{aligned}$$

By Lemma 4.7 the second term goes to zero and the second factor of the first term goes to one, leaving just

$$\begin{aligned} &= \lim_{n \rightarrow \infty} P(G_{\ell,n} \text{ is abelian} \mid D < \log \ell) \\ &= \frac{1}{e^{2c^2}} \end{aligned}$$

by Lemma 4.8. \blacksquare

4.3 Part 3 of Theorem 1.1: when $\ell \in \omega(\sqrt{n})$ and $\ell \in o(n)$.

By Lemma 4.3 we know that when $\ell \in o(n)$ supercommuting is asymptotically the same as commuting. Therefore to show that asymptotically $G_{\ell,n}$ is almost never abelian we only need to show that V and W almost never supercommute. To show this, we consider $n-1$ "bins", one for each A_i . We think of each element V_i as a ball of a particular type, say red. Similarly each of the elements W_i correspond to a blue ball. We throw the ℓ red balls, and ℓ blue balls into the $n-1$ bins, and look for a particular collision that implies V and W don't supercommute. To prove this, we will use the following Lemma which is a generalized (to 2 colors) version of the probabilistic pigeonhole principle. A statement for q -colors appears in [3].

Fact 4.10 (Lemma 5 in [3]). Let μ be any probability measure on a set of size n . Let $z_1, \dots, z_{2\ell}$ be chosen randomly and independently using μ . Then

$$P(\exists i, j \text{ with } i \leq \ell < j, z_i = z_j) \geq 1 - 2e^{-c\ell/\sqrt{n}}$$

for some universal constant c . □

In particular, when $\ell \in \omega(\sqrt{n})$, this probability approaches 1 as $n \rightarrow \infty$.

Lemma 4.11. When $\ell \in \omega(\sqrt{n})$ as $n \rightarrow \infty$ the probability that V, W supercommute goes to zero. □

Proof. Let f be the function that takes $A_k^{\pm 1}$ to k , and define 2ℓ random variables $\{z_i\}$ as follows: when $i \leq \ell$,

$$z_i = f(V_i)$$

and when $i > \ell$,

$$z_i = \begin{cases} n-1 & \text{if } f(W_{i-\ell}) = 1 \\ f(W_{i-\ell}) - 1 & \text{otherwise} \end{cases}$$

Then the conditions of Fact 4.10 apply to the z_i 's, and so asymptotically almost surely there exist an i and j so that $i \leq \ell < j$ and $z_i = z_j$. This means that either $z_i = f(V_i) = f(W_{j-\ell}) - 1 = z_j$ or $f(V_i) = n-1$ and $f(W_{j-\ell}) = 1$. The latter case has probability $1/(n-1)$, and so as $n \rightarrow \infty$ we are almost surely in the former case. Thus $V_i = A_k^{\pm 1}$ and $W_{j-\ell} = A_{k+1}^{\pm 1}$. Then V_i and W_j do not commute, and so V and W do not supercommute. ■

Corollary 4.12. If $\ell = \omega(\sqrt{n})$ and $\ell = o(n)$ then $G_{\ell, n}$ is asymptotically almost surely nonabelian. □

Proof. By Lemma 4.11 the probability that V, W supercommute goes to zero and therefore by Lemma 4.3, $G_{\ell, n}$ is asymptotically almost surely nonabelian. ■

4.4 Part 3 of Theorem 1.1: when $\ell \in \omega(n)$

In this case we need results from Section 3 on the distribution of superdiagonal entries.

Lemma 4.13. When $\ell \in \omega(n)$ then $G_{\ell, n}$ is a.a.s. not abelian. □

Proof. By Lemma 2.1, if $v_1 w_2 \neq v_2 w_1$ then $G_{\ell, n}$ is not abelian. By Lemma 3.1, $P(w_2 = 0) \sim K\sqrt{n/\ell} \rightarrow 0$. Then by Lemma 3.6, $P(v_1 w_2 = v_2 w_1) = P(v_1 w_2 - v_2 w_1 = 0) \rightarrow 0$, and so a.a.s. $v_1 w_2 \neq v_2 w_1$. ■

4.5 Part 3 of Theorem 1.1: when $k \leq \ell/n \leq M$

To complete the proof of Theorem 1.1 part 3, we need to consider functions ℓ which lie in the complement of $o(n)$, and $\omega(n)$; we therefore consider functions ℓ such that for large enough n , there exists constants k and M so that

$$k \leq \frac{\ell}{n} \leq M.$$

To show that $G_{\ell, n}$ is not abelian, it is sufficient to find $1 \leq i \leq n-2$ for which the condition of Lemma 2.1 fails; that is, there exists an i so that $v_i w_{i+1} \neq v_{i+1} w_i$. To do this, we count a subset of pairs of words V and W which have this property, and show that these pairs occur with high probability.

Lemma 4.14. Suppose there exist constants k and M so that for large enough n , $k \leq \ell/n \leq M$. Then a.a.s. there is some $1 \leq i \leq n-2$ for which $v_i = \pm 1$, $v_{i+1} = 0$, $w_i = \pm 1$, and $w_{i+1} = \pm 1$. □

Proof. We will look specifically for cases in which $V_j = A_i^{\pm 1}$ for precisely one j , $V_j \neq A_{i+1}^{\pm 1}$ for all j , $W_j = A_i^{\pm 1}$ for precisely one j , and $W_j = A_{i+1}^{\pm 1}$ for precisely one j . Note that words V and W of this form have $v_i = \pm 1$, $v_{i+1} = 0$, $w_i = \pm 1$, and $w_{i+1} = \pm 1$. Hence, by Lemma 2.1, V and W will not commute. It'll be useful to have a name for this sort of failure to commute, so we'll say this particular sort of pair (V, W) has a "type i " configuration. Our strategy for this proof is to define a random variable X which counts the expected number of type i configurations for a pair of words (V, W) . We then show $E[X^2]/E[X]^2 \rightarrow 1$. It will be sufficient to consider only odd values of i , and as this makes some of the counting arguments simpler, we make this assumption.

Fix i . Let S_i be the set of words V of length ℓ which have $V_j = A_i^{\pm 1}$ for precisely one j and $V_j \neq A_{i+1}^{\pm 1}$ for all j . There are ℓ indices to choose for the location of $A_i^{\pm 1}$, two choices for the exponent on A_i , and after subtracting out the elements $A_i^{\pm 1}$ and $A_{i+1}^{\pm 1}$, we have $2(n-3)$ remaining generators to choose from for the remaining $\ell-1$ elements in the word V . Since the total number of words of length ℓ is $(2(n-1))^\ell$, we have

$$P(S_i) = \frac{\ell(n-3)^{\ell-1}}{(n-1)^\ell} = \frac{\ell}{n-1} \frac{(n-3)^{\ell-1}}{(n-1)^{\ell-1}} = \frac{\ell}{n-1} \left(1 - \frac{2}{n-1}\right)^{\ell-1}.$$

Let T_i be the set of words W for which $W_j = A_i^{\pm 1}$ and exactly one j' for which $W_{j'} = A_{i+1}^{\pm 1}$. Then we have

$$P(T_i) = \frac{\ell(\ell-1)(n-3)^{\ell-2}}{(n-1)^\ell} = \frac{\ell(\ell-1)}{(n-1)^2} \left(1 - \frac{2}{n-1}\right)^{\ell-2}.$$

Since V and W are chosen independently, we have

$$P(S_i, T_i) = \frac{\ell^2(\ell-1)}{(n-1)^3} \left(1 - \frac{2}{n-1}\right)^{2\ell-3}. \quad (4)$$

Now we compute the probability of $S_i \cap S_{i'}$, for distinct i and i' . Counting words of this sort is where we use the convenience of only considering odd indices, so that $|i-i'| \geq 2$.

$$P(S_i \cap S_{i'}) = \frac{\ell(\ell-1)(n-5)^{\ell-2}}{(n-1)^\ell} = \frac{\ell(\ell-1)}{(n-1)^2} \left(1 - \frac{4}{n-1}\right)^{\ell-2}.$$

Similarly, we compute the probability of $T_i \cap T_{i'}$.

$$P(T_i \cap T_{i'}) = \frac{\ell(\ell-1)(\ell-2)(\ell-3)(n-5)^{\ell-4}}{(n-1)^\ell} = \frac{\ell(\ell-1)(\ell-2)(\ell-3)}{(n-1)^4} \left(1 - \frac{4}{n-1}\right)^{\ell-4}.$$

Let n' be the number of odd integers in $[1, n-2]$, and let X be the number of odd values of i for which a type i configuration occurs in the pair (V, W) . Define the random variable X_i

$$X_i = \begin{cases} 1 & \text{if } V \text{ is in } S_i \text{ and } W \text{ in } T_i \\ 0 & \text{otherwise} \end{cases}$$

Then $X = \sum_{i=1}^{n'} X_{2i-1}$ and

$$E(X) = n'P(S_i, T_i).$$

Note that when ℓ is in the complement of $o(n)$, we have $E(X) \rightarrow \infty$ as $n \rightarrow \infty$. (Also, when ℓ is in $\omega(n)$, the expected value $E(x) \rightarrow 0$ as $n \rightarrow \infty$, hence this proof is not valid when ℓ is in this range.)

When $i \neq i'$, $X_i X_{i'} = 1$ if and only if V is in $S_i \cap S_{i'}$ and W is in $T_i \cap T_{i'}$. Therefore,

$$E(X^2) = n'P(S_i, T_i) + n'(n'-1)P(S_i \cap S_{i'}, T_i \cap T_{i'}).$$

We now argue that $E(X^2)/E(X)^2 \rightarrow 1$ as $n \rightarrow \infty$.

$$\begin{aligned} \frac{E(X^2)}{E(X)^2} &= \frac{n'P(S_i, T_i) + n'(n'-1)P(S_i \cap S_{i'}, T_i \cap T_{i'})}{(n')^2 P(S_i, T_i)^2} \\ &= \frac{1}{n'P(S_i, T_i)} + \left(\frac{n'-1}{n'}\right) \frac{P(S_i \cap S_{i'}, T_i \cap T_{i'})}{P(S_i, T_i)^2}. \end{aligned}$$

When ℓ is bounded above by Mn , the first term goes to zero as $n \rightarrow \infty$. After simplifying a bit, we have,

$$\frac{P(S_i \cap S_{i'}, T_i \cap T_{i'})}{P(S_i, T_i)^2} = \frac{(\ell-2)(\ell-3)}{\ell^2} \left(\frac{n-5}{n-3}\right)^{2\ell-6} \left(\frac{n-3}{n-1}\right)^{-2\ell}$$

When $\ell = cn$, the product of the later two functions limits to 1. When can therefore conclude that $E[X^2]/E[X]^2 \rightarrow 1$ whenever ℓ is (eventually) bounded below by kn and above by Mn . Since $E[X] \rightarrow \infty$, asymptotically almost surely $X > 0$, meaning that there is some odd i for which a type i configuration occurs. \blacksquare

Corollary 4.15. Suppose there exists constants k and M so that for large enough n , $k \leq \ell/n \leq M$; then a.a.s. $G_{\ell, n}$ is not abelian. \square

5 Full Step

To analyze whether our group $G_{\ell,n}$ has full step we rely heavily on the results from Section 3. Define two families of indicator random variables δ and γ as follows:

$$\delta_{v,i} = \begin{cases} 1 & \text{if } v_i = 0 \\ 0 & \text{if } v_i \neq 0 \end{cases} \quad \delta_{w,i} = \begin{cases} 1 & \text{if } w_i = 0 \\ 0 & \text{if } w_i \neq 0 \end{cases} \quad \gamma_i = \begin{cases} 1 & \text{if } v_i = w_i = 0 \\ 0 & \text{if } v_i \neq 0 \text{ or } w_i \neq 0 \end{cases}$$

Note that $\gamma_i = \delta_{v,i}\delta_{w,i}$.

5.1 Part 1 of Theorem 1.2: when $\ell \in o(n^2)$

In this case we show that $G_{\ell,n}$ is a.a.s never full step but we separate the proofs into two subcases. In Corollary 5.2 we consider the case when $\ell \in O(n)$ while in Lemma 5.3 we consider the case when $\ell \in \omega(n) \cap o(n^2)$. The following lemma is standard but is the basis for Corollary 5.2 so we include the proof.

Lemma 5.1. If cn balls are thrown uniformly and independently into n bins, there is a.a.s. at least one empty bin. \square

Proof. Let X be the number of empty bins. Then $X = \sum_i X_i$ where

$$X_i = \begin{cases} 1 & \text{if bin } i \text{ is empty} \\ 0 & \text{otherwise.} \end{cases}$$

Then

$$\begin{aligned} \mathbb{E}(X) &= n \mathbb{E}(X_i) \\ &= n \mathbb{P}(\text{Bin } i \text{ is empty}) \\ &= n \left(1 - \frac{1}{n}\right)^{cn} \end{aligned}$$

and

$$\begin{aligned} \mathbb{E}(X^2) &= \mathbb{E}(X) + 2 \sum_{i \neq j} \mathbb{E}(X_i X_j) \\ &= n \left(1 - \frac{1}{n}\right)^{cn} + 2 \frac{n(n-1)}{2} \mathbb{P}(\text{Bins } i \text{ and } j \text{ are both empty}) \\ &= n \left(1 - \frac{1}{n}\right)^{cn} + n(n-1) \left(1 - \frac{2}{n}\right)^{cn}. \end{aligned}$$

Thus $\mathbb{E}(X) \rightarrow \infty$ and

$$\begin{aligned} \frac{\mathbb{E}(X^2)}{\mathbb{E}(X)^2} &= \frac{n \left(1 - \frac{1}{n}\right)^{cn} + n(n-1) \left(1 - \frac{2}{n}\right)^{cn}}{n^2 \left(1 - \frac{1}{n}\right)^{2cn}} \\ &\sim \frac{\left(1 - \frac{2}{n}\right)^{cn}}{\left(1 - \frac{1}{n}\right)^{2cn}} \\ &\rightarrow 1, \end{aligned}$$

And so $\mathbb{P}(X = 0) \rightarrow 0$. Thus a.a.s. $X > 0$, and so there is at least one empty bin. \blacksquare

Corollary 5.2. If $\ell \in O(n)$, a.a.s. $G_{\ell,n}$ is not full-step. \square

Proof. Let $V = V_1 \cdots V_\ell$ and $W = W_1 \cdots W_\ell$. Set up $n - 1$ bins and put a ball in bin i whenever some $V_j = A_i^{\pm 1}$ or $W_j = A_i^{\pm 1}$. Note that this process effectively throws in 2ℓ balls uniformly and independently into the $n - 1$ bins. Since $\ell \in O(n)$, there is some $c > 0$ for which $2\ell < c(n - 1)$ for large enough n , and thus by Lemma 5.1 there is an empty bin. This empty bin corresponds to some i for which $v_i = w_i = 0$, and so by Lemma 2.4 $G_{\ell,n}$ is not full-step. \blacksquare

Lemma 5.3. If $\ell \in o(n^2)$ and $\ell \in \omega(n)$, a.a.s. $G_{\ell,n}$ is not full-step. \square

Proof. Let X be the number of positions on the superdiagonal for which V and W both have a 0. That is

$$\begin{aligned} X &= \sum_i \gamma_i \\ \mathbb{E}(X) &= \sum_i \mathbb{E}(\gamma_i) \\ &= n \mathbb{P}(v_i = w_i = 0). \end{aligned}$$

By Corollary 3.2,

$$\begin{aligned} &\sim nK^2 \frac{n}{\ell} \\ &\sim K^2 \frac{n^2}{\ell} \\ &\rightarrow \infty \end{aligned}$$

when $\ell \in o(n^2)$.

Also,

$$\begin{aligned} \mathbb{E}(X^2) &= \mathbb{E} \left[\left(\sum_i \gamma_i \right)^2 \right] \\ &= \sum_i \mathbb{E}(\gamma_i) + 2 \sum_{i \neq j} \mathbb{E}(\gamma_i \gamma_j) \\ &= \sum_i \mathbb{P}(v_i = w_i = 0) + 2 \sum_{i \neq j} \mathbb{P}(v_i = v_j = w_i = w_j = 0). \end{aligned}$$

By Corollaries 3.2 and 3.5,

$$\begin{aligned} &\sim nK^2 \frac{n}{\ell} + n^2 K^4 \frac{n^2}{\ell^2} \\ &= K^2 \frac{n^2}{\ell} + K^4 \frac{n^4}{\ell^2}. \end{aligned}$$

Then

$$\frac{\mathbb{E}(X^2)}{\mathbb{E}(X)^2} \sim \frac{K^2 \frac{n^2}{\ell} + K^4 \frac{n^4}{\ell^2}}{K^4 \frac{n^4}{\ell^2}}$$

and since $\ell \in o(n^2)$ the second term dominates in the numerator to give us

$$\begin{aligned} &\sim \frac{K^4 \frac{n^4}{\ell^2}}{K^4 \frac{n^4}{\ell^2}} \\ &\sim 1. \end{aligned}$$

Since $\mathbb{E}(X) \rightarrow \infty$ and $\mathbb{E}(X^2)/\mathbb{E}(X)^2 \rightarrow 1$ then $\mathbb{P}(X > 0) \rightarrow 1$. So there is at least one i for which $\gamma_i = 1$, that is $v_i = w_i = 0$. Then by Lemma 2.4 we have that $G_{\ell,n}$ is not full-step. \blacksquare

5.2 Part 2 of Theorem 1.2: when $\ell \in \omega(n^3)$

Lemma 5.4. If $\ell \in \omega(n^3)$, a.a.s. $G_{\ell,n}$ is full-step. \square

Proof. Let X be the number of zeroes on the superdiagonal of W . That is

$$X = \sum_i \delta_{w,i}.$$

Then

$$E(X) = \sum_i E(\delta_{w,i}).$$

Since the δ are identically distributed,

$$= nP(w_i = 0).$$

By Lemma 3.1,

$$\begin{aligned} &\sim nK\sqrt{\frac{n}{\ell}} \\ &\sim K\sqrt{\frac{n^3}{\ell}} \\ &\rightarrow 0 \end{aligned}$$

when $\ell \in \omega(n^3)$. This means that $P(X = 0) \rightarrow 1$, and so a.a.s. none of the w_i are 0.

Now, for $G_{\ell,n}$ to be full-step (that is, step $n - 1$), the $(n - 2)$ -commutator subgroup must have a nontrivial element. In particular, consider the commutator

$$C^{n-2} = \underbrace{[W, [W, \dots [W, V]]]}_{n-2}.$$

As we saw in Example 2.8 in Section 2 the upper-right corner entry of C^{n-2} is given by

$$c_{n,n}^{n-2} = K_1 v_1 w_2 w_3 \cdots w_{n-1} + K_2 w_1 v_2 w_3 \cdots w_{n-1} + \cdots + K_{n-1} w_1 \cdots w_{n-2} v_{n-1}$$

where each $K_i = \binom{n-1}{i}$ with alternating signs. Since the w_i and K_i are a.a.s. nonzero and $\ell \in \omega(n)$, Lemma 3.6 says that $P(c_{n,n}^{n-2} = 0) \rightarrow 0$ and thus a.a.s. $c_{n,n}^{n-2} \neq 0$, making C^{n-2} nontrivial. ■

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